

# Credit Supply and Firms' Productivity Growth

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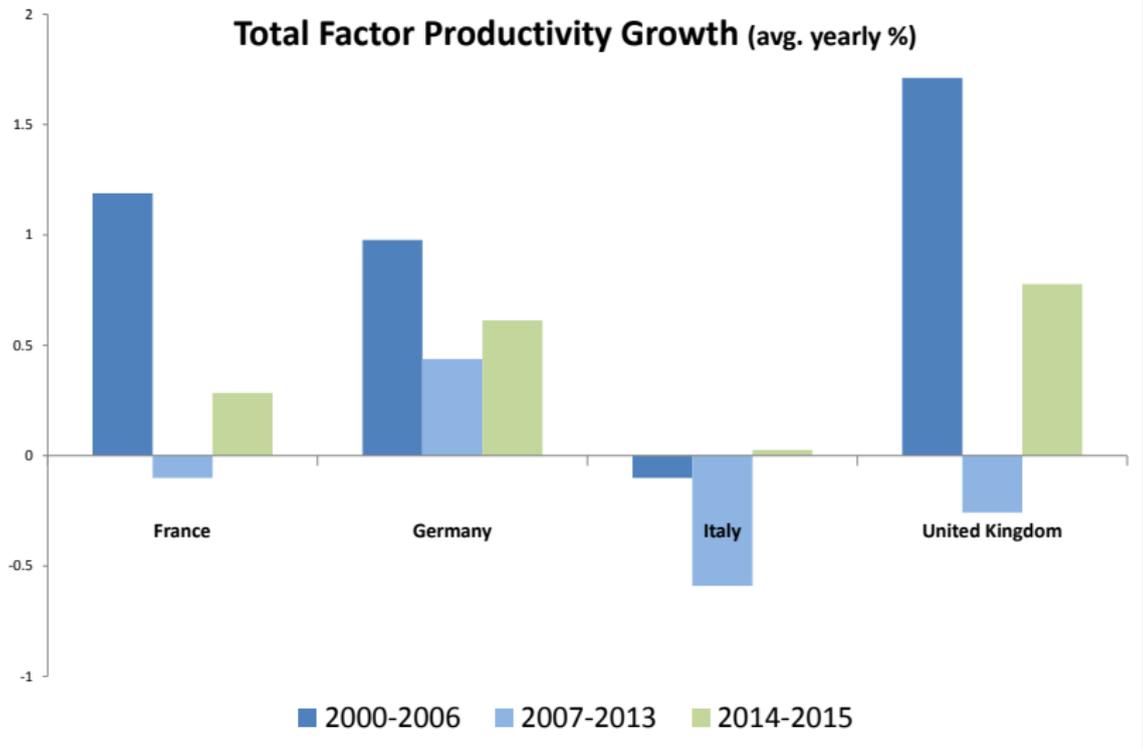
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# Introduction

- ▶ The financial and sovereign crises have witnessed significant TFP slowdown in Europe;
- ▶ Growth afterwards remained sluggish.

### Total Factor Productivity Growth (avg. yearly %)



# Introduction

- ▶ Several explanations for recent TFP trends:
  - ▶ “secular” stagnation
  - ▶ faltering innovation
  - ▶ slowdown in business dynamism
  - ▶ output data fail to capture values of new digital products

Does credit supply play a role?

## Research Question

Olley-Pakes decomposition of average productivity:

$$\sum_i \omega_{i,t} \cdot marketshare_{i,t} = \frac{1}{N} \sum_i \omega_{i,t} + cov(\tilde{\omega}_{i,t}, marketshare_{i,t})$$

- ▶ **Credit Supply and TFP via Input Misallocation:** Midrigan and Xu (2014), Gopinath et al. (2016).
- ▶ Yet, there are reasons to expect also a **direct effect on firm-level TFP growth**  $\omega_{i,t}$ :  
through innovation (Amore et al. 2013), export (Paravisini et al. 2014), technology adoption, managerial practices.

# The impact of credit supply shocks in the literature

Growing literature on identification of firm-level credit supply shocks from firm-bank matched data (Khwaja-Mian, Greenstone-Mas-Nguyen, Amiti-Weinstein)

- ▶ **Labor:** Chodorow-Reich (2014), Bentolilla et al. (2016).
- ▶ **Investments:** Gan (2007) Cingano et al. (2016), Acharya et al. (2016); Bottero et al. (2016).
  
- ▶ So far, no study in this literature plugged results into a production function framework.
  
- ▶ Some contemporaneous papers on “credit constraints  $\Rightarrow$  TFP” (Dorr et al.; Duval et al.; de Sousa and Ottaviano)

# This paper

- 1. Identifying firm-level changes in credit supply:**
  - ▶ exploits bank-firm matched data + stickiness of lending relationships
- 2. Estimates TFP allowing for an effect of credit supply on TFP**
  - ▶ productivity process allowed to be directly affected by credit supply
- 3. Estimates the effect of credit supply on TFP**
  - ▶ **main results:**  $\uparrow 1\%$  cred supply  $\Rightarrow \uparrow 0.13\%$  productivity growth [▶ More Results](#)
  - ▶ BotE calculation: a drop in credit growth of around 12 p.p. (2006-2008)  $\Rightarrow 25\%$  aggregate reduction in TFP over the same period
  - ▶ persistent effect on productivity *levels*
- 4. Beyond measurement: channels**
  - ▶ evidence that credit supply boosts Export & Innovation (R&D and Patenting)

# Data

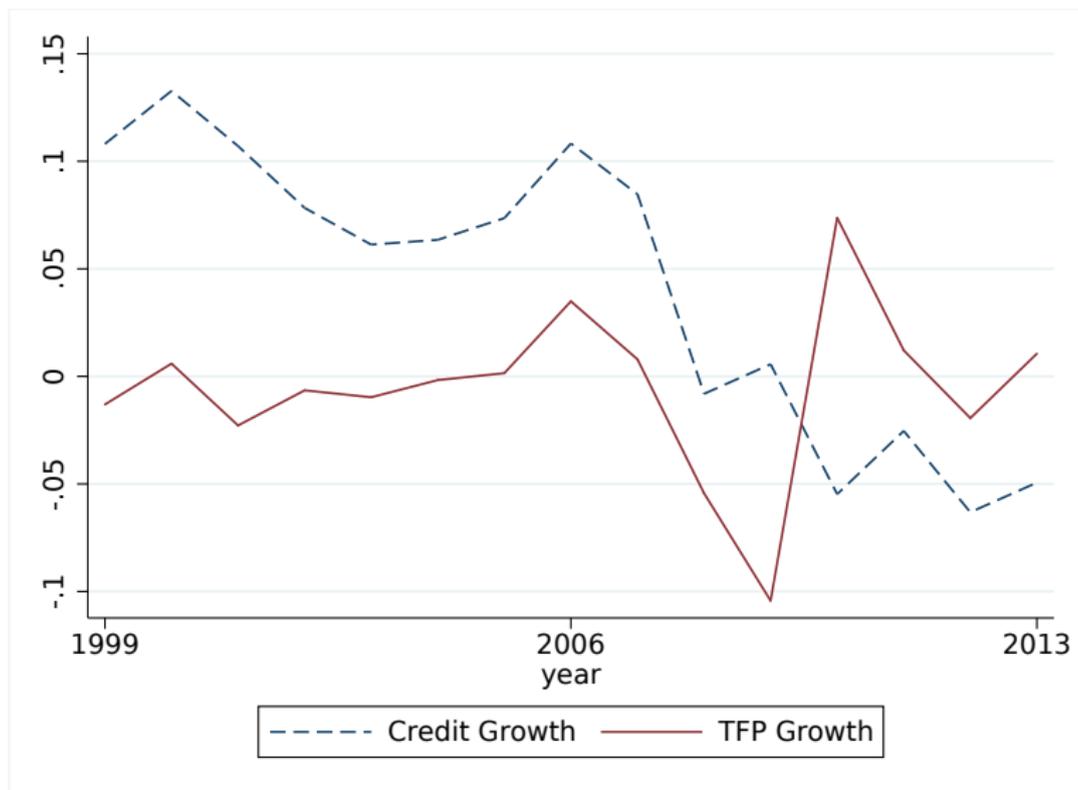
Credit Register: all credit relations in country

- ▶ report credit instruments, we use total
- ▶ focus on credit granted, yearly
- ▶ on average, per year:
  - ▶ 468,984 firms
  - ▶ 1,008 banks
  - ▶ 2.8 relationships per firm; 1,321 per bank

Balance-Sheets and Income Statement from CADS:

- ▶ large sample of small and large Italian manufacturers
- ▶ capital series reconstructed with perpetual inventory methodology
- ▶ sector-level deflators from National Accounts
  - ▶  $\Rightarrow$  measure of productivity based on revenues, not quantity (Foster, Haltiwanger, and Syverson, 2008)

## TFP and Credit among Italian Firms



Notes: Data from ~ 30K Italian firms from CADS dataset.

Estimated p.f.: Value Added Cobb-Douglas.

# Identifying Credit Supply Shocks

## Credit Supply: an Empirical Framework

Total credit granted to firm  $i$  at the end of year  $t$  is equal to

$$C_{i,t} = \sum_b C_{ib,t}$$

Assume, as a starting point:

$$\frac{C_{ib,t}}{C_{ib,t-1}} = \frac{C(\delta_t, U_{i,t}, U_{b,t})}{C(\delta_{t-1}, U_{i,t-1}, U_{b,t-1})}$$

Log-linearizing:

$$\Delta C_{ibt} = c_t + \Delta U_{it} + \Delta U_{bt} + \varepsilon_{ibt}$$

## A Valid Decomposition?

We are assuming away:

- ▶ **no assortative matching:** firm demand is not bank-specific
- ▶ **no granularity in credit demand:** firms are sufficiently small
- ▶ **no spillover across banks** because of substitutability/complementarities btw banks.

$$\Delta c_{ibt} = c_t + \Delta u_{it} + \Delta u_{bt} + \Delta u_{ibt} + d\Delta u_{b't} + \varepsilon_{ibt}$$

# Tackling Identification Assumptions

- ▶ **assortative matching:** test robustness of results against controlling for firm-bank (lagged) characteristics
  - ▶ length of lending relationship
  - ▶ share of collateral
  - ▶ share of drawn credit
  - ▶ interest rate charged
- ▶ **granularity in credit demand:** exclude top-borrowers
- ▶ **spillover across banks:** iterating procedure → include supply shocks of other banks (main bank or avg lenders) previously estimated

Resulting estimates of  $\Delta u_{bt}$  are very similar ( $Corr. \approx .90$ ).  
All results on productivity confirmed.

# From Bank Shocks to Firm Credit Supply

- ▶ We compute firm-level credit supply shocks as:

$$\chi_{it} = \sum_b w_{ib,t-1} \Delta u_{bt}$$

where  $w_{ib,t-1} = C_{ib,t-1} / C_{i,t-1}$

- ▶ Logic of  $w_{ibt}$ : Borrower-lender relations mitigate asymmetric info & limited commitment
  - ▶ valuables, costly to establish and sticky
  - ▶  $\Rightarrow$  changes in lenders' credit supply affects financing ability of connected borrowers

# Measuring Productivity

## A simple theoretical model

Production function:

$$Y_{i,t} = \exp\{\omega_{i,t} + \epsilon_{i,t}^Y\} F(L_{i,t}, K_{i,t}, M_{i,t}, \beta)$$

s.t.

$$K_{i,t} = I_{i,t} + (1 - \delta_t)K_{i,t-1}$$

$$\tilde{\pi}_{i,t} + B_{i,t} = D_{it} + I_{i,t} + B_{i,t-1} (1 + r_{i,t}) + \textit{Adjustment}$$

$$B_{i,t} \leq K_{i,t-1} \cdot \Gamma(\chi_{i,t}, \omega_{i,t})$$

Taking logs:

$$y_{i,t} = \omega_{i,t} + \epsilon_{i,t}^Y + f(k_{i,t}, l_{i,t}, m_{i,t}, \beta)$$

Assuming intermediates  $m_{i,t}$  are fully flexible and monotonic is monotonic in  $\omega_{i,t}$ , we invert its demand function

$$\omega_{i,t} = m^{-1}(m_{i,t}, k_{i,t}, l_{i,t}, z_{i,t}, \chi_{i,t}) \Rightarrow$$

$$y_{i,t} = m^{-1}(m_{i,t}, k_{i,t}, l_{i,t}, z_{i,t}, \chi_{i,t}) + f(k_{i,t}, l_{i,t}, m_{i,t}, \beta) + \epsilon_{i,t}^Y \Rightarrow$$

First stage estimation:

$$y_{i,t} = \Psi(m_{i,t}, k_{i,t}, l_{i,t}, z_{i,t}, \chi_{i,t}) + \epsilon_{i,t}^Y$$

## Productivity law of motion

$$E[\omega_{i,t} | \mathcal{I}_{t-1}] = g_t(\omega_{t-1}, \chi_{i,t-1})$$

approximate  $g$  with a polynomial

$$\zeta_{i,t} := \omega_{i,t} - g(\omega_{t-1}, \chi_{i,t-1})$$

$$\Rightarrow E[\zeta_{i,t} | \mathcal{I}_{t-1}] = 0$$

what does it mean?

1.  $\nexists$  persistent, firm-specific unobservable affecting input choices and productivity
  - ▶ violated if we did not include  $\chi_{i,t}$ .
2. shocks to  $\omega$  are orthogonal to lagged variables
  - ▶ violated if e.g. company invested more in the past anticipating higher prod growth

## Estimating moments

$$E[\zeta_{i,t} + \xi_{i,t} | \mathcal{I}_{t-1}] = 0 \Rightarrow$$

$$E \left[ \begin{array}{c} y_{i,t} - f(k_{i,t}, l_{i,t}, m_{i,t}, \beta) - g(\Psi_{i,t-1} - f(k_{i,t-1}, l_{i,t-1}, m_{i,t}, \beta), \chi_{i,t-1}, G_t) \\ \Psi_{i,t-1} \\ m_{i,t-1} \\ \dots \end{array} \middle| \begin{array}{c} l_{i,t-1} \\ inv_{i,t-1} \\ \Psi_{i,t-1} \\ m_{i,t-1} \\ \dots \end{array} \right]$$

$\Rightarrow$  estimate, for each industry, both  $\beta$  and the ancillary coefficients  $G_t$

- ▶ value added: average  $\beta_k \approx 0.35$  and  $\beta_l \approx 0.64$
- ▶ net revenues: average  $\beta_k \approx 0.03$ ,  $\beta_l \approx 0.10$  and  $\beta_m \approx 0.87$

# Results

# Credit supply and input & output growth

For each (log) input or output measure we estimate:

$$\Delta x_{i,t} = \psi_i + \psi_{p,s,t} + \gamma x_{i,t} + \eta_{i,t}$$

VARIABLES	(1) $\Delta va$	(2) $\Delta y$	(3) $\Delta k$	(4) $\Delta l$	(5) $\Delta n$	(6) $\Delta m$
$x_{i,t}$	0.144*** (0.0227)	0.0477*** (0.0158)	0.0572*** (0.0192)	-0.0271 (0.0184)	-0.0126 (0.0127)	0.0126 (0.0167)
Observations	293k	293k	293k	293k	293k	293k
R-squared	0.248	0.320	0.260	0.258	0.324	0.319

▶ All sectors

# Credit supply and productivity

Now we can run:

$$\Delta\omega_{i,t} = a_i + \psi_{p,s,t} + \gamma\chi_{i,t} + \eta_{i,t}$$

	(1)	(2)	(3)	(4)
$\chi_{i,t}$	0.133*** (0.0241)	0.138*** (0.0205)	0.0423*** (0.00820)	0.0510*** (0.00749)
Output measure	va	va	y	y
Functional Form	CD	TL	CD	TL
Observations	278k	258k	286k	272k
R-squared	0.198	0.339	0.159	0.271

► All sectors

## On the effect of credit supply on productivity

- ▶ Results show that the effect is significant and positive: a 1 p.p. increase in credit supply triggers VA productivity by 0.13 p.p.
- ▶ Results less different between VA and revenues productivity, once effects are standardized.
- ▶ Effect stronger for smaller firms, and in manufacturing.

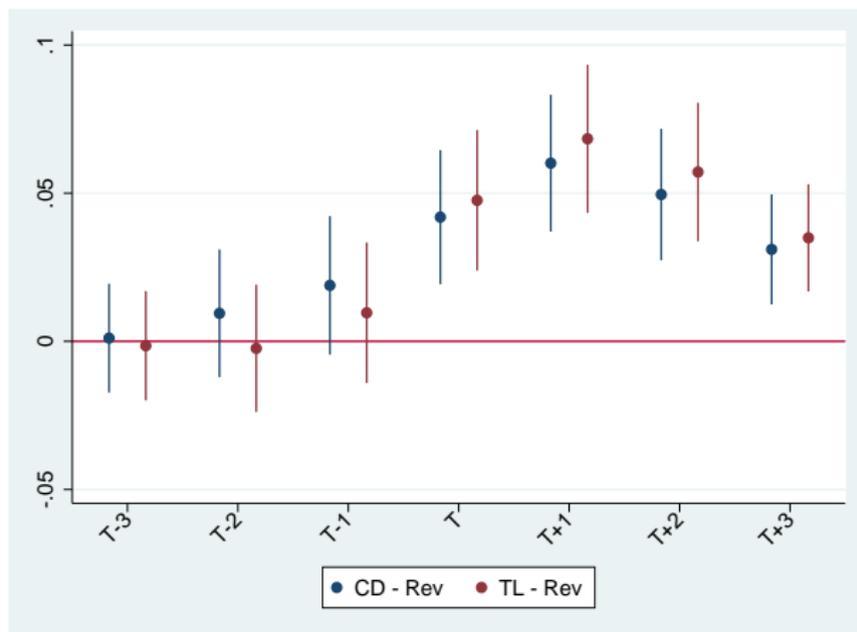
## Estimated effect is remarkably robust

Results are unaffected by

- ▶ inclusion of firm-level controls;
- ▶ use of different Fixed-Effects structure (test for correlated unobservables);
- ▶ estimate of bank shocks net of spillovers & controlling for assortative matching btw firms and banks;
- ▶ exclusion of top-3% (“granular”) borrowers;
- ▶ controlling for impact of credit supply on firm’s demand ⇒ firms involved into global and local VC are NOT differently affected [▶ Results](#)
  
- ▶ use of a different identification strategy for credit supply shocks: the 2007-2008 collapse of the interbank mkt.

# Persistency and Pre-trend

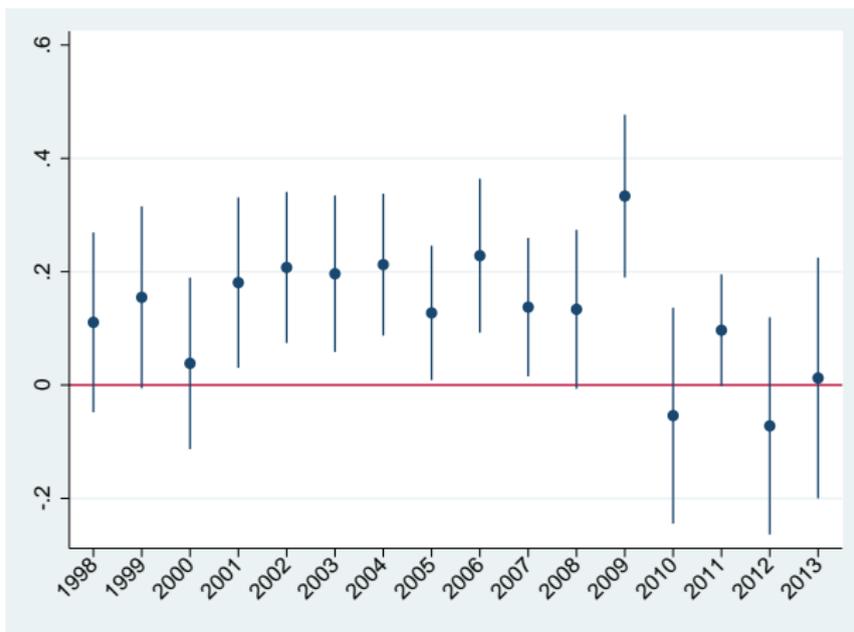
$$\omega_{i,t} = \psi_i + \psi_{p,s,t} + \sum_{j=3}^{-3} \gamma_j \chi_{i,t-j} + \eta_{i,t}$$



No significant pre-trend, levels remain persistently higher after shock.

## Effects over time

$$\Delta\omega_{i,t} = \psi_i + \psi_{p,s,t} + \sum_t \gamma_t \chi^i_{i,t} + \eta_{i,t}$$



Effect peaks in 2009, but significant also before crisis

# **Why Does Credit Availability Enhance Productivity Growth?**

# Additional Data

## INVIND

- ▶ survey conducted from '84 on panel of firms
- ▶ mostly >50 employees
- ▶ some waves have info on innovation and export activities
- ▶ neither questions nor respondents are fixed over time

## Patents

- ▶ Patents registered at EPO by all Italian firms;
- ▶ Matched to fiscal codes by the Italian Chamber of Commerce (Unioncamere);
- ▶ Priority Dates : 1999-2012.

## Possible Mechanisms? ICT adoption

Number of PC used by the firm available for years 1999, 2000, 2001

- ▶ do firms become more ICT intense when credit constraints are more relax

$$\log \left( \frac{PC}{employees} \right)_{i,t} = \gamma_i + \gamma_t + \alpha \chi_{i,t} + \eta_{i,t}$$

and

$$\log \left( \frac{PC}{K} \right)_{i,t} = \gamma_i + \gamma_t + \alpha \chi_{i,t} + \eta_{i,t}$$

## Results

No statistically significant evidence of positive effect

VARIABLES	(1) $\log\left(\frac{PCs}{employees}\right)$	(2) $\log\left(\frac{PCs}{employees}\right)$	(3) $\log\left(\frac{PCs}{K}\right)$	(4) $\log\left(\frac{PCs}{K}\right)$
$\chi_{i,t}$	0.117 (0.149)	0.302 (0.282)	0.257 (0.220)	0.513 (0.379)
Obs	6541	1969	6232	2193
Sample	All	Exclude top 25%	All	Exclude top 25%
$R^2$	0.935	0.932	0.939	0.921

Clustered standard errors in parentheses

Firm and year FE are included

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

## Possible Mechanisms? - R&D and Export

High quality information on size of R&D investment from INVIND

- ▶ we consider dicotomic variables
  - ▶ exporter vs non-exporter (dummy  $Expt_{i,t}$ )
  - ▶ positive versus zero R&D investment
- ▶ we have two measures of R&D
  - ▶  $R\&D_{i,t}$
  - ▶  $RD\&E_{i,t}$

LPM with firm fixed effect:

$$Pr(d_{i,t} = 1) = \gamma_i + \gamma_t + \alpha\chi_{i,t} + \eta_{i,t}$$

where  $d_{i,t}$  is any of the dummies described above

## Results

Companies are more likely to start (less to stop) exporting or doing R&D (only one of our measures) when credit availability increases

VARIABLES	(1) <i>Expt<sub>i,t</sub></i>	(2) <i>R&amp;D<sub>i,t</sub></i>	(3) <i>RD&amp;E<sub>tal,i,t</sub></i>
$\chi_{i,t}$	0.152* (0.085)	0.238* (0.128)	-0.064 (0.105)
Obs	13,249	5,991	15,177

Clustered standard errors in parentheses

Firm and year FE are included

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

## Possible Mechanisms? Financial constraints to innovation

Innovative effort is much broader than just formal R&D or ITC adoption

- ▶ 2011 survey wave investigate which were the main constraints to innovative effort for previous year
- ▶ one question ask how important were difficulties to collect external funds in limiting innovation on a four-items scale
- ▶  $FinCon_{i,2010}$  equal to one iff difficulties to get external funds is thought to be “somehow important” or “very important” as obstacle to innovation

## Result - Financial constraints to innovation

Linear Probability Model, using cross section

$$Pr(\text{FinCon}_{i,2010} = 1) = \gamma_{s,p} + \alpha\chi_{i,2010} + \eta_{i,t}$$

Estimates

- ▶  $\hat{\alpha} = -1.111^*$
- ▶  $tstat = -1.75$
- ▶  $N=628$
- ▶ caveats: only regression with  $\chi_{i,t}$  without firm FE (we include *province*  $\times$  *sector*)

$\Rightarrow$  Innovation efforts are less likely to be constraints by lack of external funds when firms just received a positive credit shock

## Possible Mechanisms? Patenting

$$\#Pat_{i,t} = a_i + \gamma_{s,p,t} + \alpha\chi_{i,t} + \varepsilon_{i,t}$$

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	(1)	(2)	(3)
$\chi_{i,t}$	0.032*** (0.010)	0.038** (0.018)	0.036** (0.017)
Firm FE	N	Y	Y
Sector FE	Y	Y	N
Province FE	Y	Y	N
Year FE	Y	Y	N
Sec-Prov-Year FE	N	N	Y
Obs	241K	241K	241K

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Clustered standard errors in parentheses

Firm and year FE are included

\*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$

# Conclusion

In this paper we

- ▶ exploit banks-firms connections to measure firm-specific shocks to credit supply
- ▶ estimate a simple model of production with heterogeneous credit frictions
- ▶ show that productivity growth is boosted by increase in credit supply
- ▶ document that productivity enhancing activities are stimulated by credit availability

What's next:

- ▶ improve our identification of possible mechanisms
- ▶ compute relative importance of credit frictions for allocative efficiency vs productivity growth

# Thank You

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# All sectors

VARIABLES	(1) $\Delta va$	(2) $\Delta y$	(3) $\Delta k$	(4) $\Delta l$	(5) $\Delta n$	(6) $\Delta m$
$\chi_{i,t}$	0.106*** (0.0182)	0.0424*** (0.0121)	0.0531*** (0.0144)	0.00461 (0.0140)	0.00144 (0.0104)	0.0233* (0.0125)
Observations	552k	552k	552k	552k	552k	551k
R-squared	0.232	0.311	0.264	0.276	0.324	0.312

VARIABLES	(1) $\Delta \omega_{i,t}$	(2) $\Delta \omega_{i,t}$	(3) $\Delta \omega_{i,t}$	(4) $\Delta \omega_{i,t}$
$\chi_{i,t}$	0.0890*** (0.0175)	0.106*** (0.0183)	0.0173*** (0.00523)	0.0244*** (0.00547)
Observations	552k	552k	551k	551k
R-squared	0.179	0.191	0.192	0.212
Output measure	va	va	revenues	revenues
Functional Form	CD	TL	CD	TL

▶ Back - Productivity

▶ Back - Inputs/Outputs

## Direct Effect on Demand

Bank might directly affect borrowers demand because of correlation between lenders of suppliers and lenders of clients (e.g. local effect). Then we run

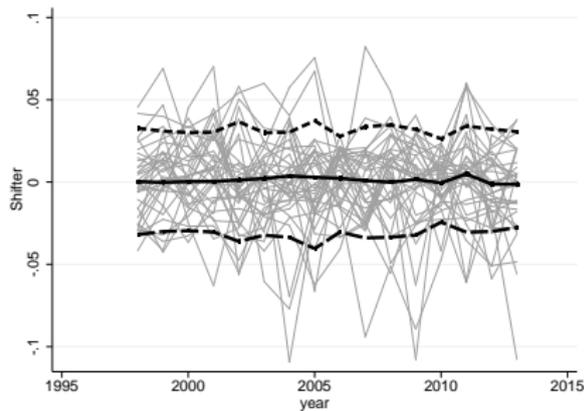
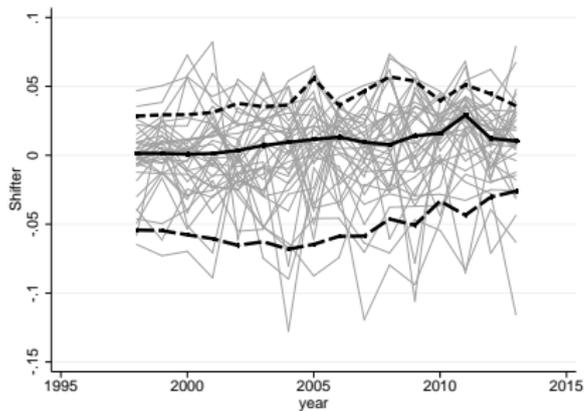
$$\Delta\omega_{i,t} = \psi_t + \psi_i + \gamma_0\chi_{i,t} + \gamma_1\frac{\text{export}_{i,t-2}}{y_{i,t-2}} + \gamma_2\chi_{i,t} \cdot \frac{\text{export}_{i,t-2}}{y_{i,t-2}} + \eta_{i,t}$$

$\gamma_2$  capture the differential effect of the shock on exporters

- ▶ less likely foreign buyers land from same bank  $\Rightarrow \gamma_2 < 0$
- ▶ results: not statistically different from zero
- ▶  $\Rightarrow$  effects does not come from direct effect on mark up

## “Visualizing” the relevant variation: RHS

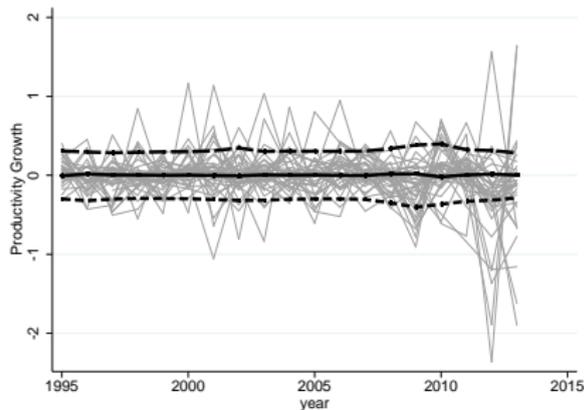
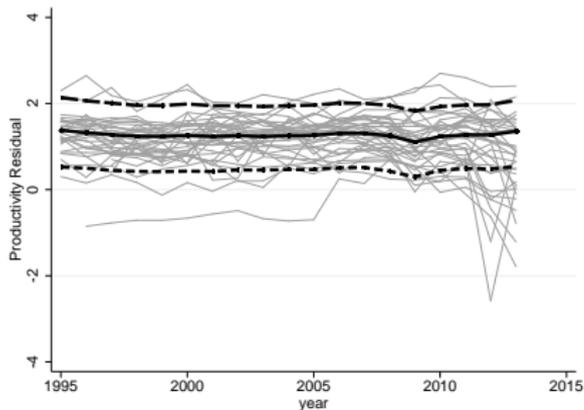
Evolution of  $\chi_{i,t}$  for a 5% random sample of manufacturers



- ▶ Right panel shows residualized values after taking out FEs
- ▶ no clear pattern over time:  $\chi_{i,t}$  makes sense only relatively

## “Visualizing” the relevant variation: LHS

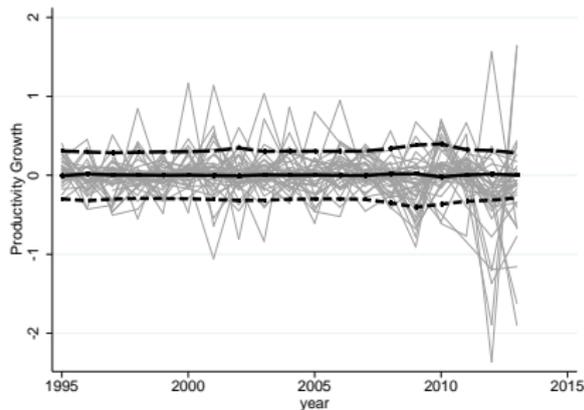
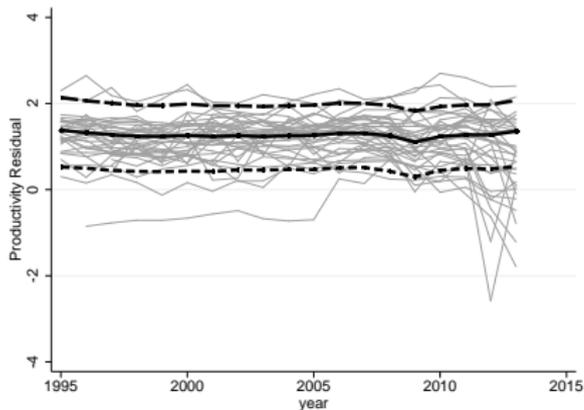
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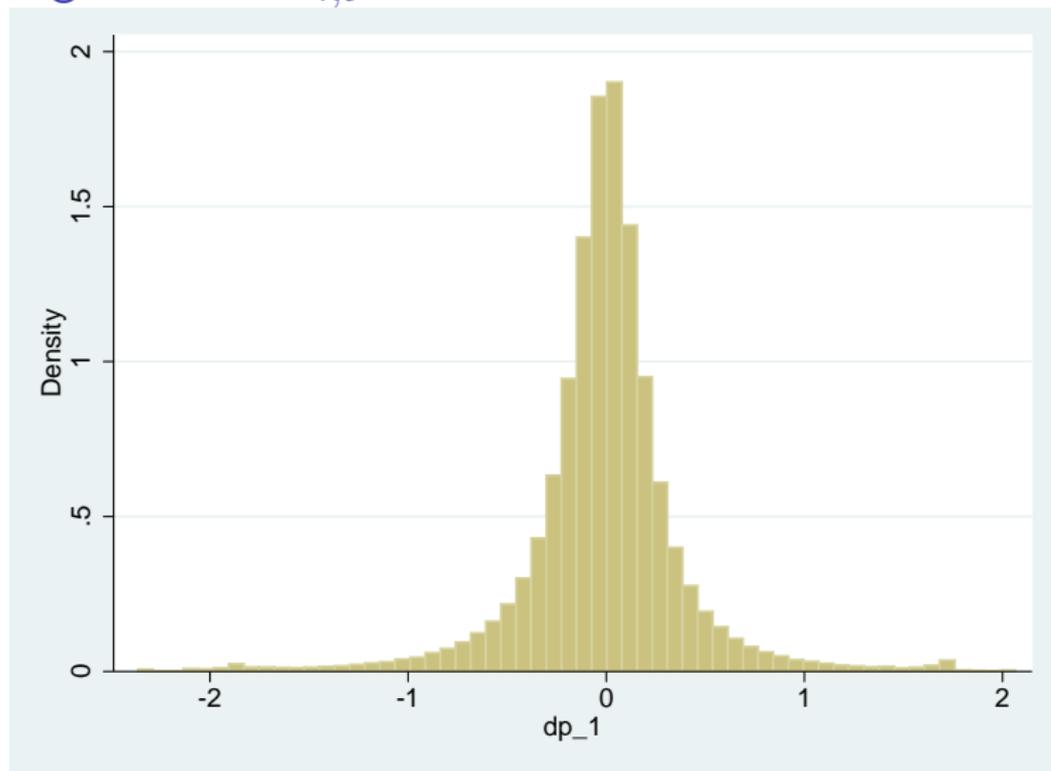
## “Visualizing” the relevant variation: LHS

Evolution of  $\omega_{i,t}$  for a 5% random sample of manufacturers



- ▶ Right panel shows residualized values after taking out FEs

## Histogram of $\Delta\omega_{i,t}$



► Value Added - Cobb Douglas